

MARKOV RANDOM WALKS ON GROUPS

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Abstract.

Let X be a Markov chain on C and f a continuous function from C to a topological group G . The process

$$S_0 = e, \quad S_{n+1} = S_n f(X_{n+1}), \quad n = 1, 2, \dots$$

is called a Markov random walk on G . We investigate the set of possible values taken by S and, under some regularity conditions on X , we also study the recurrence/transience properties of S by means of an embedded ordinary random walk on G . We show, for instance, that if G is a transient group then all Markov random walks on G are transient.

Let G be a multiplicatively written locally compact second countable topological group. The object of this paper is to investigate the sequence of successive products S_1, S_2, \dots of factors of the form $f(X_1), f(X_2), \dots$ where the X_n 's form a Markov chain on some topological space C and f is a continuous function from C to G . In particular, we are interested in the possible values and the recurrence properties of the sequence $(S_n)_{n=0}^\infty$.

Muthsam [8] and Wolff [13] studied the problem for the case G a discrete semigroup and G a compact semigroup, respectively. (In their work $C \subseteq G$ and S_n is simply $X_1 X_2 \dots X_n$.) We will draw heavily on these two papers.

Niemi and Nummelin [9] studied central limit properties of $(S_n)_{n=0}^\infty$, with $G = \mathbb{R}$ and f an arbitrary measurable function. In our investigation, we will make use of their technique of introducing an artificial recurrent atom for the Markov chain.

1.

Let C be a locally compact second countable Hausdorff space and $(X_n)_{n=0}^\infty$ a Markov chain on C with transition kernel P , which we will assume to be Feller, cf. [12, p. 34]. P_x will denote the probability on the

canonical probability space (path space) C^∞ induced by P and the initial condition $X_0 = x$. Let f be a continuous function from C to G . The process

$$S_n = \prod_{k=1}^n f(X_k), \quad n = 1, 2, \dots; \quad S_0 = e$$

will be called a *Markov random walk* on G .

EXAMPLE 1. Let X be a Markov chain on the integers \mathbb{Z} , with transition probability matrix P . Then

$$S_n = X_1 + X_2 + \dots + X_n, \quad n = 1, 2, \dots$$

is a Markov random walk on \mathbb{Z} . It reduces to ordinary random walk if the rows of P are all identical (i. e. the X_i 's are independent). If $C = \{1, -1\}$ with $P(1,1) = P(-1,1) = \frac{1}{2}$, then S is a symmetric simple random walk.

EXAMPLE 2. Let $C = \{0\} \cup \{n^{-1} \mid 1, 2, \dots\}$ endowed with the topology of the real line. Let P be defined as follows: $P(0,1) = 1$

$$P(n^{-1}, (n+1)^{-1}) = \alpha_n, \quad P(n^{-1}, 1) = 1 - \alpha_n, \quad n = 1, 2, \dots$$

where $0 < \alpha_n < 1$, $n = 1, 2, \dots$. If X is the random walk on C with the transition probability matrix P , then

$$S_n = X_1 + X_2 + \dots + X_n, \quad n = 1, 2, \dots$$

is a Markov random walk on the real line \mathbb{R} .

EXAMPLE 3. Let μ be a probability distribution on the set of idempotent matrices of the form

$$(1 + \alpha\beta)^{-1} \begin{pmatrix} 1 & \alpha \\ \beta & \alpha\beta \end{pmatrix}$$

where $\alpha, \beta \in \mathbb{R}$, $\alpha\beta \neq -1$. The behavior of a product

$$X_1 X_2 \dots X_{n-1} X_n$$

of independent μ -distributed random matrices is conveniently (for most purposes) studied by considering the (1,1) entry of the product:

$$S_n = (1 + \alpha_1\beta_1)^{-1} (1 + \alpha_1\beta_2) \dots (1 + \alpha_{n-1}\beta_n)(1 + \alpha_n\beta_n)^{-1}$$

where the subscript i corresponds to the random matrix X_i . As we shall see below S_n is, under mild conditions on the probability μ , a Markov random walk on the multiplicative group $\mathbb{R} \setminus \{0\}$.

In our analysis of the Markov random walks we will make extensive use of the fact that the auxiliary process $Z_n = (X_n, S_n)$, $n = 0, 1, 2, \dots$, is a Markov chain on $C \times G$ with transition kernel

$$U((x, s), A \times B) = \int 1_A(z) 1_B(sf(z)) P(x, dz)$$

$$x \in C, s \in G, A \in \mathcal{C}, B \in \mathcal{G}$$

where \mathcal{C} and \mathcal{G} are Borel σ -algebras of C and G , respectively.

Following [13] we call a $(n + 1)$ -tuple (x_0, x_1, \dots, x_n) a *chain* of length n from x to y if $x = x_0$, $y = x_n$, and x_{i+1} is in the support of the measure $P(x_i, \cdot)$, $i = 0, 1, \dots, n - 1$. Denote the chain by k_x and define the functions p and q by

$$p(k_x) = f(x_1) f(x_2) \dots f(x_n) \quad \text{and}$$

$$q(k_x) = x_n.$$

Suppose a sequence $(k_x^n)_{n=1}^\infty$ of chains has the property that $q(k_x^n)$ converges to u . If w is a limit point of the sequence $p(k_x^n)$, it is necessarily of the form $w = vf(u)$, since $p(k_x^n) = v_n f(q(k_x^n))$ for some $v_n \in G$. Note that the assumptions that G is a group and f a continuous function are used in this argument.

For $x, u \in G$ consider the set of all sequences of chains (k_x^n) starting from x such that $q(k_x^n)$ converges to u (the sequence may be finite if there is a chain from x to u).

DEFINITION. $\Pi(x, u)$ is the set of all limit points of the sequences $p(k_x^n)$ such that $q(k_x^n)$ converges to u .

If $q(k_x^n) \rightarrow u$, then we say that x *leads to* u , $x \rightsquigarrow u$. If $x \rightsquigarrow u$ implies $u \rightsquigarrow x$, we say that x is *essential*. Note that $x \rightsquigarrow u$ is equivalent to

$$u \in \text{closure} \left(\bigcup_{n=1}^\infty \text{supp}(P^n(x, \cdot)) \right),$$

where P^n is the n -step transition kernel for X .

We will make the following general assumption about the Markov chain X (cf. [13]),

- (A) C is an irreducible class and all elements of C are essential.
- (A) guarantees that all elements of C communicate. I.e. for every $x, u \in G$ there exists a sequence (k_x^n) of chains starting from x such that $q(k_x^n)$ converges to u .

In our Example 1 above $\Pi(0, 0)$ is the set of all sums $x_1 + x_2 + \dots + x_{n-1} + 0$ such that the transition probabilities $P(0, x_1), P(x_1, x_2), \dots,$

$P(x_{n-1}, 0)$ are all positive. If, for example, $P(0, nd), P(nd, 0), P(0, -md), P(-md, 0) > 0$ for some relatively prime positive integers $m, n,$ and $d,$ then $\Pi(0, 0) \cong d\mathbb{Z}.$

In Example 2, $\Pi(1, 1)$ is a subset of the real interval $[1, \infty).$ Note that $\Pi(1, 1)$ is necessarily a semigroup and that it is asymptotically dense. It is also easily seen that $\Pi(0, 1) = \Pi(1, 1)$ but $\Pi(0, 0)$ is empty because all chains with $q(k_0^n)$ close to 0 have

$$p(k_0^n) \geq 1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{m_n}$$

for some large positive integer $m_n.$

$\Pi(x, x)$ is either empty or a closed subsemigroup of $G,$ the *invariance semigroup* of the element $x \in C.$ To see this, notice that (x, g) can be reached by the Z -chain from $(x, e),$ if and only if $g \in \Pi(x, x),$ cf. [13]. So if $(x, e) \rightsquigarrow (x, h)$ ((x, e) leads to (x, h)) as well we get $(x, e) \rightsquigarrow (x, g)$ and $(x, g) \rightsquigarrow (x, gh)$ whence, by the transitivity of the relation $\rightsquigarrow,$ we obtain the desired result $(x, e) \rightsquigarrow (x, gh).$

REMARK. The intuitively obvious transitivity of \rightsquigarrow can be seen as follows, cf. [13]: If g is a bounded measurable function on $C \times G,$ then we can define

$$Ug(x, s) = \int g(z, sf(z)) P(x, dz) \quad (x, s) \in C \times G$$

$(U((x, s), N))$ is to be interpreted as $U1_N(x, s)$ for a set $N \in \mathcal{C} \times \mathcal{G}.$

Using the assumption that P is Feller, the local compactness of C and the continuity of the multiplication in G we can show that U and all its iterates U^n are Feller. Note that the functions $U^n(\cdot, N)$ are then lower semicontinuous for open sets $N.$ Suppose $(x, s) \rightsquigarrow (x', s')$ and $(x', s') \rightsquigarrow (x'', s'').$ This means, by definition, that for any neighborhoods N' and N'' of (x', s') and $(x'', s''),$ respectively, there are n' and n'' with $U^{n'}((x, s), N'), U^{n''}((x', s'), N'') > 0.$ Choose N' to be such that $U^{n''}(\cdot, N'') > 0$ on N' to obtain $U^{n'+n''}((x, s), N'') > 0.$

Suppose that there is an essential element $(c, g) \in C \times G$ for the Z -chain. Clearly, $(c, g) \rightsquigarrow (c, g)$ so $e \in \Pi(c, c).$ If $h \in \Pi(c, c),$ then $(c, g) \rightsquigarrow (c, gh)$ and, by essentiality, $(c, gh) \rightsquigarrow (c, g),$ that is $h^{-1} \in \Pi(c, c).$ All elements communicating with (c, g) are essential (as well as all $(c, h), h \in G$) so $\Pi(u, u)$ is also a group for all those $u \in C$ for which $\Pi(c, u) \neq \emptyset.$ This is, for instance, the case for all

$$u \in \bigcup_{n \geq 1} \text{support}(P^n(c, \cdot)),$$

a set which is dense in C by assumption (A). The set of essential elements thus contains the set

$$\bigcup_{\Pi(c,u) \neq \emptyset} \{u\} \times G,$$

cf. [8], [13].

The groups $\Pi(c,c)$ and $\Pi(u,u)$ are, in fact, conjugate. To see this, consider the transitions $(u,e) \rightsquigarrow (c,l) \rightsquigarrow (u,e)$ or, equivalently, $(u,e) \rightsquigarrow (c,l)$ and $(c,e) \rightsquigarrow (u,l^{-1})$. If $g \in \Pi(c,c)$, we get $lgl^{-1} \in \Pi(u,u)$ and, conversely, every element of $\Pi(u,u)$ can be represented in this way. Hence $l\Pi(c,c)l^{-1} = \Pi(u,u)$.

REMARK. A slight modification of Example 2 will make $\Pi(1,1)$ a group (= R): Introduce a negative element -2 , say, and a transition probability $P(1, -2) > 0$ and let $P(-2, 1) = 1$. Then $\Pi(x,y) = \mathbb{R}$ for all $x,y \in C$. Note that the assumption (B) below is not satisfied by the Markov chain in Example 2 even in its modified form.

For any (Feller) Markov chain, the essential classes are closed. If, for example, the group G is compact or $\Pi(c,c) = G$, we can then conclude that all elements of $C \times G$ are essential. This also holds, of course, for the case when

$$(B) \quad C = \bigcup_{n \geq 1} \text{support}(P^n(c, \cdot)).$$

Conversely, if (B) is true for all $c \in C$ we can argue as in [8] to prove that $\Pi(c,c)$ is a group (for one and thus for all $c \in C$) if and only if (c,g) is essential for the Z -chain (for some $g \in G$).

We summarize the preceding discussion in

THEOREM 1. *Under assumption (A), if there is an essential element (c,g) for the Z -chain the set of essential elements contains the set*

$$\bigcup_{\Pi(c,u) \neq \emptyset} \{u\} \times G.$$

The invariance semigroup $\Pi(c,c)$ is a closed subgroup of G . Furthermore, for u such that $\Pi(c,u) \neq \emptyset$, $\Pi(u,u)$ is a group conjugate to $\Pi(c,c)$.

COROLLARY 1. *If, in addition, condition (B) holds for all $c \in C$, either all or none of the invariance semigroups $\Pi(c,c)$ are groups. Then all or none, respectively, of the elements of $C \times G$ are essential for the Z -chain.*

COROLLARY 2. *If $\Pi(c,c) = G$ or if G is compact, then the same holds for any $u \in C$ and all elements of $C \times G$ are essential.*

For compact C and G we refer to the thorough discussion in [13].

REMARK. It might be worth while to note that the (relative) compactness of $\Pi(c,c)$ will automatically make it a group, since a compact subsemigroup of a group is, in fact, a subgroup.

2.

If X is *stationary*, i.e. X has an invariant probability measure π as its initial distribution, then S is a special case of what we could call a *generalized random walk* on G , cf. [3], [7]. For a stationary X on a countable state space to admit a unique invariant probability distribution it is necessary and sufficient that X be ergodic, cf. [13, p. 136]. In the case of a general state space we will use the sufficient condition that X be positive Harris, see [12]. In this section, we will study the *recurrence* properties of such a walk. For the case $G = \mathbb{R}$ the criterion for recurrence is particularly simple [3]:

S is recurrent if and only if $E_\pi f(X_1) (= E_\pi S_1) = 0$

(E_x is expectation with respect to P_x and $E_\pi(\cdot) = \int E_x(\cdot \pi(dx))$)

If G is the direct product of a compact group K and \mathbb{R} , we have a similar criterion, namely $E_\pi f_2(X_1) = 0$, where f_2 is the second component of f : If $f(x) = (k,r) \in K \times \mathbb{R}$, then $f_2(x) = r$. This result follows from the preceding, since the walk on the compact factor is automatically recurrent and the walk on the second factor \mathbb{R} is itself nothing but a generalized random walk on \mathbb{R} . For $G = \mathbb{C}^*$, the multiplicative group of non-zero complex numbers $\simeq T \times \mathbb{R}$, where T is the circle group, we thus have the following recurrence criterion

$$E_\pi \log |S_1| = 0.$$

Consider a group G . If all (ordinary) random walks generated by probability measures whose supports are not contained in any proper closed subgroups of G are transient, the group itself is called a *transient group*. It is well-known that \mathbb{R}^m is transient for $m \geq 3$. Under mild conditions Markov random walks are also transient on \mathbb{R}^m , $m \geq 3$, see [4]. Our aim is to show that the same holds true for all transient groups provided that some additional assumptions are satisfied.

Introduce the following condition on the transition kernel P of the Markov chain X , cf. [9, (1.7)]

$$(1) \quad P(x,A) \geq h(x)v(A), \quad x \in C, \quad A \in \mathcal{C},$$

where ν is a measure on (C, \mathcal{C}) and $h \geq 0$ a function on C with $\int h(x)\pi(dx) > 0$.

Using the splitting technique described in [9], cf. also [1], we can decompose the chain $X = (X_1, X_2, \dots)$ by (a.s. finite) random times $\tau(1), \tau(2), \dots$ in such a way that the blocks $X_{\tau(i-1)+1}, \dots, X_{\tau(i)}$ and $X_{\tau(i)+1}, \dots, X_{\tau(i+1)}$ are independent and identically distributed (i.i.d.). (For a thorough exposition of the technique the reader should consult the fundamental [10]).

REMARK. For a Harris recurrent chain, P^k is minorized as in (1) for some positive integer k , see [9]. The choice of $k = 1$ is necessary in order to ensure the decomposition of X_1, X_2, \dots into *independent* blocks.

Suppose first that C is discrete and that X is positive recurrent and aperiodic. (1) is automatically satisfied with $h(x) = 1_{\{c\}}(x)$, $\nu = P(c, \cdot)$, where c is an arbitrary fixed element of C . The times $\tau(1), \tau(2), \dots$ are the successive returns of X to c . They are all a.s. finite: $E_\pi \tau(1) = (\pi(c))^{-1}$.

$$X_{\tau(1)+1} X_{\tau(1)+2} \dots X_{\tau(n)} = S_{\tau(1)}^{-1} S_{\tau(n)} \quad n = 1, 2, \dots$$

are products of i.i.d. random elements of $\Pi(c, c) \cong G$. In other words, $(S_{\tau(n)}^{-1} S_{\tau(n)})_{n=1}^\infty$ is a random walk on $\Pi(c, c)$, generating all of $\Pi(c, c)$. Hence if $\Pi(c, c)$ is a transient group the random walk is transient. If $\Pi(c, c)$ is a subsemigroup of G which is not a group, then the random walk must be transient, too (because any recurrent random walk automatically generates a group, see [12]). The same holds for the process $(S_{\tau(n)})_{n=1}^\infty$, which differs from $(S_{\tau(1)}^{-1} S_{\tau(n)})_{n=1}^\infty$ only by its initial distribution. Now, by (an extension to the group case of) the Main Lemma of [1], we can conclude that the Markov random walk $(S_n)_{n=1}^\infty$ itself is transient. Hence the desired result – $\Pi(c, c)$ transient or no group at all $\Rightarrow S$ transient – holds for discrete C . (From the discussion in Section 1 we know that the invariance semigroups $\Pi(c, c)$ are all isomorphic if they are groups).

In the general case the times $\tau(i)$ are return times to the set, where the function h is positive, but, in general, not all of them. The difficulty lies in the problem of characterizing the set (subgroup), where the random walk $(S_{\tau(i)}^{-1} S_{\tau(n)})_{n=1}^\infty$ “lives”, i.e. the smallest closed subgroup containing all the random variables $S_{\tau(1)}^{-1} S_{\tau(n)}$, $n = 1, 2, \dots$. Under certain conditions on f and ν , a characterization of this subgroup can be obtained. Suppose, for example, that the support of ν has non-empty interior and that the function f is an open mapping. Then the smallest closed subgroup that could possibly contain $f(\text{supp}(\nu))$ is the identity component of G . If G is connected and transient it follows that S is transient.

We can formulate this sufficient condition for transience slightly more

generally. The distribution of the random group element $f(Z)$, where Z is distributed according to ν , is a measure νf^{-1} on the Borel sets of G . If the support of νf^{-1} generates all of G , then $(S_{\tau(1)}^{-1} S_{\tau(n)})_{n=1}^{\infty}$ lives on G proper, i. e. there is no closed subgroup a. s. containing all the products. Hence this random walk is transient if G is a transient group. As before, it follows that $(S_n)_{n=1}^{\infty}$ is transient, too.

We summarize our results in a

THEOREM 2. *Let X be positive Harris satisfying (1). If*

- (i) *C is discrete and the invariance semigroups $\Pi(c, c)$, $c \in C$, are either no groups at all or transient groups or*
 - (ii) *the smallest closed subgroup containing the support of the measure νf^{-1} ($\nu f^{-1}(A) = \nu(f^{-1}(A))$, $A \in \mathcal{G}$) is G itself and G is transient,*
- then S is transient.*

We saw above that for $G = \mathbb{R}$ or $K \times \mathbb{R}$ the stationarity of X is enough to furnish us with a recurrence criterion, namely $E_{\pi} S_1 = 0$ and $E_{\pi} f_2(X) = 0$, respectively. For other groups, we will again use the decomposition of X_1, X_2, \dots into independent blocks. If $(S_{\tau(1)}^{-1} S_{\tau(n)})_{n=1}^{\infty}$ is recurrent, then a fortiori $(S_n)_{n=1}^{\infty}$ is.

A random walk S on \mathbb{R}^2 is recurrent if and only if $ES_1 = 0$ and $ES_1^2 < \infty$, cf. [12, p. 98]. Hence our Markov random walk on \mathbb{R}^2 will be recurrent if and only if

$$E_{\pi}(S_{\tau(2)} - S_{\tau(1)}) = 0 \quad \text{and} \quad E_{\pi}|S_{\tau(2)} - S_{\tau(1)}|^2 < \infty$$

(assuming, as before, that the process does not live on a smaller subgroup). The first condition reduces to $E_{\pi} S_1 = 0$, since $E_{\pi} S_{\tau} = (E_{\pi} \tau)(E_{\pi} S_1)$ (with $\tau = \tau(1)$). The second condition may be written

$$E_{\nu} \left| \sum_{i=0}^{\tau} f(X_i) \right|^2 < \infty.$$

If f is bounded, $E_{\nu} \tau^2 < \infty$ will guarantee the recurrence of S .

THEOREM 3. *Let X be as in Theorem 2. If*

- (i) *C is discrete, the invariance semigroups $\Pi(c, c) = \mathbb{R}$ or \mathbb{R}^2 and $E_{\pi} S_1 = 0$ and, in addition, in the \mathbb{R}^2 -case,*

$$E_c \left| \sum_{i=1}^{\tau} f(X_i) \right|^2 < \infty \quad \text{for some } c \in C,$$

where τ is the first return time to c , or

- (ii) $G = \mathbb{R}$ and $E_\pi S_1 = 0$ or
- (iii) $G = \mathbb{R}^2$, $E_\pi S_1 = 0$ and

$$E_\nu \left| \sum_{i=0}^{\tau(1)} f(X_i) \right|^2 < \infty,$$

where $\tau(1), \tau(2), \dots$ are the times discussed above decomposing the chain X_1, X_2, \dots into i.i.d. blocks,

then the Markov random walk S is recurrent.

The conditions given are also necessary (in (ii) and (iii) provided that the distribution of $S_{\tau(2)} - S_{\tau(1)}$ is not supported by a proper closed subgroup).

REMARK. If G is compact (or $\Pi(c, c)$ in the discrete case) then S is, of course, always recurrent.

3.

For all Harris chains condition (1) is valid for some P^k . However, it is crucial for the decomposition of X into i.i.d. blocks that $k = 1$. This limitation is a considerable drawback, when we want to investigate Markov random walks, where the increment $S_{n+1} - S_n$ is a function of two (or more) consecutive values, e.g., $S_{n+1} - S_n = f(X_n, X_{n+1})$.

The process S is still a Markov random walk: just consider $\hat{X}_n = (X_n, X_{n+1}) \in C \times C$, cf. [11], but property (1) is lost in general.

$$P(\hat{X}_{n+1} \in A \times B \mid \hat{X}_n = (x_1, x_2)) = \delta_{x_2}(A)P(X_{n+2} \in B \mid X_{n+1} = x_2) \geq \delta_{x_2}(A)h(x_2)\nu(B)$$

which cannot be written in the form (1) unless, of course, $\pi\{x_2\} > 0$. The two-step transition is again more regular,

$$P(\hat{X}_{n+2} \in A \times B \mid \hat{X}_n = (x_1, x_2)) \geq h(x_2) \left(\int_A \nu(dy)h(y) \right) \nu(B).$$

Random walks on certain groups, notably the group of rigid motions on \mathbb{R}^d , have been investigated using the theory of semi-Markov processes, cf. [4]. It can also be regarded as a Markov random walk on \mathbb{R}^d . The essential features of random walks on the semigroup of real $n \times n$ -matrices of rank $k < n$ may be studied by viewing (an embedded) Markov random walk on the general linear group $GL(k, \mathbb{R})$, see [5] or the survey [6].

Our Example 3 is a simple example of such a walk, the product of

random projections of \mathbb{R}^2 on 1-dimensional subspaces. Almost all of these projections can be represented as matrices of the form

$$\begin{pmatrix} (1 + \alpha\beta)^{-1} & (1 + \alpha\beta)^{-1}\alpha \\ \beta(1 + \alpha\beta)^{-1} & \beta(1 + \alpha\beta)^{-1}\alpha \end{pmatrix}$$

where $\alpha, \beta \in \mathbb{R}$, $\alpha\beta \neq -1$. The projection is orthogonal if its nullspace, spanned by the vector $(-\alpha, 1)$, and its range, spanned by $(1, \beta)$, are orthogonal; thus the criterion for orthogonality is $\alpha = \beta$.

The product of two projections P_1 and P_2 is not, in general, a projection itself (below, the subscripts i refer to the matrix P_i , $i = 1, 2$):

$$P_1 P_2 = (1 + \alpha_1 \beta_1)^{-1} (1 + \alpha_1 \beta_2) (1 + \alpha_2 \beta_2)^{-1} \begin{pmatrix} 1 & \alpha_2 \\ \beta_1 & \beta_1 \alpha_2 \end{pmatrix}.$$

Suppose $X = (X_1, X_2, \dots)$ is a stationary Markov chain on the set of matrices of the same type as P above. The first element of the product matrix $X_1 X_2 \dots X_n$ is a Markov random walk on the multiplicative semigroup \mathbb{R} . Call that element S_n , then

$$S_n = \prod_{k=1}^n f(X_{k-1}, X_k), \quad S_0 = 1, \quad S_1 = (1 + \alpha_1 \beta_1)^{-1},$$

where

$$f(X_{k-1}, X_k) = (1 + \alpha_{k-1} \beta_k) (1 + \alpha_k \beta_k)^{-1}, \quad k = 2, 3, \dots$$

Assume that $P_\pi(f(X_k, X_{k+1}) = 0) = 0$ for all k (where π is the invariant probability measure for the chain X). Then S is a Markov random walk on the multiplicative group $\mathbb{R} \setminus \{0\}$. Hence it is recurrent if and only if

$$\begin{aligned} E_\pi \log |f(X_0, X_1)| \\ = \int \log |f(x, y)| \pi(dx) P(x, dy) = 0. \end{aligned}$$

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